

# Evidence of syntactic working memory usage in MEG data

**Marten van Schijndel**  
Department of Linguistics  
The Ohio State University  
vanschm@ling.osu.edu

**Brian Murphy\***  
School of Electronics,  
Electrical Engineering and Computer Science  
Queen's University Belfast  
brian.murphy@qub.ac.uk

**William Schuler**  
Department of Linguistics  
The Ohio State University  
schuler.77@osu.edu

## Abstract

While reading times are often used to measure working memory load, frequency effects (such as surprisal or  $n$ -gram frequencies) also have strong confounding effects on reading times. This work uses a naturalistic audio corpus with magnetoencephalographic (MEG) annotations to measure working memory load during sentence processing. Alpha oscillations in posterior regions of the brain have been found to correlate with working memory load in non-linguistic tasks (Jensen et al., 2002), and the present study extends these findings to working memory load caused by syntactic center embeddings. Moreover, this work finds that frequency effects in naturally-occurring stimuli do not significantly contribute to neural oscillations in any frequency band, which suggests that many modeling claims could be tested on this sort of data even without controlling for frequency effects.

## 1 Introduction

Current accounts of sentence processing (e.g., Gibson, 2000; Lewis and Vasishth, 2005) usually involve working memory: parts of sentences are stored while unrelated material is processed, then retrieved when they can be integrated. But evidence for the role of memory in sentence processing usually comes in the form of latency measurements in self-paced reading or eye-tracking data, in which frequency effects are a powerful potential confound (Hale, 2001; Levy, 2008; Demberg and

Keller, 2008; Roark et al., 2009; Smith and Levy, 2013; van Schijndel et al., 2014). For example, the direction of the correlation between memory load and reading times has been shown to be highly sensitive to complex frequency effects (Vasishth and Lewis, 2006; Schuler and van Schijndel, 2014).

Experiments described in this paper therefore attempt to find a clearer measure of variable memory usage in sentence processing, independent of frequency influences. In particular, this paper focuses on the coherence of oscillatory neural activity between anterior and posterior areas of left cortex. Areas including the left inferior frontal gyrus and the posterior left temporal cortex have been implicated in language use, especially passive listening tasks (Hagoort and Indefrey, 2014). Synchronous firing among neurons in disparate parts of the brain is thought to be a possible mechanism for the formation of cued associations in memory by causing rapidly repeated communication between association cue neurons and association target neurons, which strengthens their connection through a process of long-term potentiation (von der Malsburg, 1995; Singer, 1999; Sederberg et al., 2003; Jensen et al., 2007; Fell and Axmacher, 2011). During periods of high memory load, synchronous firing in the alpha band is thought to be associated with inhibition of memory formation so as to protect existing cues from interference (Jensen et al., 2002; Jensen et al., 2007). If this is correct, we should expect to find high alpha power and coherence among brain regions responsible for language use when language users are processing center embedded text (e.g., the bracketed text in ‘The reporter [the senator

---

\* Formerly of the Dept. of Machine Learning, Carnegie Mellon University

met] left’). Magnetoencephalographic (MEG) imaging results reported in this paper show that this does indeed seem to be the case. Exploratory analyses with the development partition of a dataset of MEG recordings of subjects listening to narrative text revealed a strong effect for memory load on alpha-band coherence between an anterior and posterior pair of left-hemisphere sensors. Follow-on validation with a larger test partition confirmed the significance of this effect. Moreover, these effects could not be explained by frequency or sentence position predictors, unlike effects on self-paced reading and eye-tracking latencies (Demberg and Keller, 2008; Roark et al., 2009; Wu et al., 2010).

The remainder of this paper is organized as follows: Section 2 provides a brief introduction to magnetoencephalography, Section 3 describes the MEG dataset used in these experiments, Section 4 describes the oscillatory coherence measure used to evaluate phase-aligned activation, Section 5 describes the center-embedding depth predictor, Section 6 describes the regression experiments and their results, and Section 7 discusses implications of these results for some open debates about hierarchic sentence processing.

## 2 MEG Background

Magnetoencephalography (MEG), like electroencephalography (EEG), is a non-invasive means to record the electrical activity of the brain, specifically the aggregate of post-synaptic potentials produced by individual neurons. MEG has certain advantages over EEG, which is the most widely used neuroimaging technique in psycholinguistics, due to its low cost, convenience and portability. While EEG’s high temporal resolution ( $\gg 100\text{Hz}$ ) makes it suitable for examining the neural processing timeline down to the level of individual words and phonemes, its spatial resolution does not compare to other techniques like fMRI (functional magnetic resonance imaging). In addition, the signals recorded with EEG (volume currents) are distorted as they pass through the skull and tissues of the head, attenuating higher frequencies, and blurring their spatial source.

MEG records magnetic fields from the same neural sources that generate the EEG-visible voltages at the scalp. As the head is transparent to mag-

netic fields, MEG signals are less noisy, have finer spatial resolution, and capture a wider range of frequencies. The EEG signal components familiar to psycholinguists (e.g., the N400 and P600) are also visible but produce different scalp distributions in MEG recordings (Pylkkänen and Marantz, 2003; Salmelin, 2007; Service et al., 2007), because of differing spatial sensitivities: EEG and MEG are more sensitive to radial and tangential sources respectively, and MEG’s higher spatial resolution means that it is not as sensitive to deep sources. And correspondingly, any magnetic coherence between two sensors can be more reliably traced to coherence between the two corresponding regions of the brain, whereas the poor spatial resolution of EEG means that coherence between sensors does not necessarily reflect coherence between the corresponding regions of the brain.

## 3 Data Collection

This study makes use of a naturalistic audio-book listening task during MEG recording. This design allows us to examine language processing in a more ecologically realistic manner (Brennan, 2010; Wehbe et al., 2014a; Wehbe et al., 2014b), as both the participant experience (reading/listening to a story for enjoyment) and author’s aim are authentic language acts.

Participants were asked to sit still in an upright position with their eyes closed, while they listened to an 80-minute excerpt of an English-language novel. The listening task was split into 8 sections of approximately 10 minutes each, and participants had the opportunity to rest between them.

The text used was the second chapter of the novel *Heart of Darkness* by Joseph Conrad, containing 628 sentences and 12,342 word tokens. The plain-text and audio book recording used were both sourced from the Gutenberg project.<sup>1</sup>

The data was recorded at 1000Hz on a 306-channel Elekta Neuromag device at the UPMC MEG Brain Mapping Center, Pittsburgh, USA. During the experiment, the audio track was recorded in parallel to enable subsequent synchronization between the brain activity and audio-book content.

<sup>1</sup><http://www.gutenberg.org/cache/epub/219/pg219.txt>;  
<http://www.gutenberg.org/ebooks/20270>

The 306 channels are distributed across 102 locations in the device helmet. Each position has a magnetometer which measures the magnitude of the magnetic flux entering or leaving the helmet at that location. The two gradiometers measure gradients in local flux (i.e. its first derivative), each in a direction perpendicular to the other.

Informed consent was obtained from 3 healthy right-handed participants, following ethical approval provided by the Institutional Review Boards of both the University of Pittsburgh, and Carnegie Mellon University.

After recording, the MEG data was preprocessed in the following way to normalize and clean the signals. The Elekta custom MaxFilter software was used to apply SSP, SSS and tSSS methods (Taulu and Hari, 2009), correcting for head motion on a run-wise basis, and removing signal components which originated outside the recording helmet and other non-brain artefacts. The EEGLab package was then used to apply a band-pass filter between 0.01–50 Hz, down-sample to 125Hz, and apply Independent Components Analysis (Delorme and Makeig, 2003). The signal time-courses and component scalp-maps were visually inspected for eye-movement and line-noise components, but none were identified.

The parallel audio recording channel was used to identify the precise sample points at which each of the 8 audio runs began and ended (these varied as participants chose to take breaks of different lengths). The eight excerpts were then spliced together to form a continuous set of MEG signals corresponding exactly to the complete audio-book time-course. This allowed us to use speech recognition forced alignment methods (MS HTK; Woodland et al., 1994) to precisely locate the onset and offset times of each auditory word. These automatically derived onset and offset times were subsequently validated by hand.

## 4 Coherence

There are a variety of measures available that reflect the connectivity between two brain regions. This study makes use of ‘spectral coherence,’ which is sensitive both to power/energy increases registered by the relevant sensors and to the degree of phase

synchronization observed by those sensors. Spectral coherence is computed with the following formula:

$$\text{coherence}(x, y) = \frac{E[S_{xy}]}{\sqrt{E[S_{xx}] \cdot E[S_{yy}]}} \quad (1)$$

where  $x$  and  $y$  are waveform signals from two sensors, and  $S_{ij}$  is the spectral density of waveforms  $i$  and  $j$ . When  $i = j$ ,  $S$  is the power spectral density of  $i$ , and when  $i \neq j$ ,  $S$  is the cross-spectral density between  $i$  and  $j$ . The expectations in the numerator and the denominator must be obtained by averaging over multiple frequency bands, multiple instances of the same frequency band in different epochs, or over both frequency bands and epochs.<sup>2</sup> The present work adopts the second approach of averaging each frequency band over multiple epochs (see Section 6 for details), which enables higher frequency resolution than if multiple frequencies had been averaged together, though it necessarily reduces the number of trials in the dataset. This work uses the MNE-python package to compute spectral coherence (Gramfort et al., 2013; Gramfort et al., 2014).<sup>3</sup>

As a measure of the correlation between two signals, coherence can be between 0 and 1. When two signals have a constant phase difference and are of the same amplitude, their coherence is 1. As either the amplitudes diverge or the phase difference changes, the coherence approaches 0.

## 5 Center Embedding Depth

This study evaluates a measure of syntactic working memory load as a predictor of MEG coherence. A canonical means of calculating syntactic working memory load is to count the number of center embeddings in a sentence. For example, the sentence in Figure 1, ‘The cart that the horse that the man bought pulled broke,’ is thought to induce greater working memory load than the same sentence without the depth 3 region: ‘The cart that the horse pulled broke,’ (Chomsky and Miller, 1963).<sup>4</sup> The increased memory load stems from an incomplete dependency (a subject lacking a predicate in the above

<sup>2</sup>If multiple instances are not averaged in Equation 1, coherence is simply 1 (Benignus, 1969).

<sup>3</sup><http://martinos.org/mne/stable/mne-python.html>

<sup>4</sup>In fact, this is an example of *self embedding*, the most difficult form of center embedding, which was chosen for ease of exposition.

<i>d1</i>	The cart	broke.
<i>d2</i>	that the horse	pulled
<i>d3</i>	that the man bought	

Figure 1: Center embeddings in ‘The cart that the horse that the man bought pulled broke.’ Each lexeme is associated with the given embedding depth on the left.

example) that must be retained in working memory until the dependency can be completed (Gibson, 2000). The load should increase every time there is a right branch from a left branch in a syntactic binary-branching tree.<sup>5</sup>

Experiments described in this paper estimate syntactic memory load when processing a particular word of a sentence as the center-embedding depth of that word, which is the number of incomplete categories maintained while processing that word using a left-corner parser (Aho and Ullman, 1972; Johnson-Laird, 1983; Abney and Johnson, 1991; Gibson, 1991; Resnik, 1992; Stabler, 1994). To obtain an accurate estimate of center-embedding depth, this study uses the van Schijndel et al. (2013) left-corner PCFG parser trained on the Penn Treebank (Marcus et al., 1993) reannotated into a Nguyen et al. (2012) generalized categorial grammar (GCG),<sup>6</sup> which makes PCFG probabilities sensitive to filler-gap propagation. This parser achieves a linguistic accuracy comparable to the Petrov and Klein (2007) parser, and the PCFG surprisal estimates it outputs using this grammar provide a state-of-the-art fit to psycholinguistic measures like self-paced reading times and eye-tracking fixation durations (van Schijndel and Schuler, 2015).

The experiments described in Section 6 run this parser on transcripts of the *Heart of Darkness* dataset described in Section 3, calculating center-embedding depth for each word epoch based on its position in the best output parse. This parser is also used to calculate PCFG surprisal as a potentially confounding predictor.

<sup>5</sup>In fact, there are conditions where a post-modifier can create a complex left-branching structure that does not cause an associated increase in memory load, but that effect is beyond the scope of this paper.

<sup>6</sup><http://sourceforge.net/projects/modelblocks/>

## 6 Methodology

In this section we describe how we establish a reliable effect of sentence embedding depth on alpha-band coherence in the MEG recordings. While our analysis is motivated by experimental results using non-linguistic stimuli (e.g., Jensen et al., 2002), we do not expect the scalp topology of EEG effects to be exactly replicated in MEG recordings, and we do not necessarily expect coherence observations during skilled behavior like sentence comprehension to exactly match observations while processing word lists. This, and the possibility of frequency-based confounds, requires an exploratory analysis of a range of sensor-pairs, frequency bands, and time windows. To avoid the danger of selection biases we partition one third of the data into a development set and the rest of the data into a test set. The development data gives an indication of which sensor pair best reflects a stable correlation between embedding depth and MEG coherence, which is later confirmed using the test partition.

The van Schijndel et al. (2013) parser is used to obtain estimates of the embedding depth of each word in the corpus according to the best output parse of the sentence. As described in Section 5, these estimates are used as a measure of the memory load that is present as each word is processed.

The data is divided into epochs, which extend from one second pre-onset to two seconds post-onset for each word. This window extends beyond the average auditory duration of a word ( $\sim 0.4s$ ), and assumes that the processing timeline for each word is time-locked to its auditory onset (Hagoort, 2008). In order to clean up extraneous noise in the signal, words are omitted if they are in a sentence that fails to parse, if they are in an extremely short or an extremely long sentence ( $<4$  or  $>50$  words), or if they follow a word at a different depth, which could introduce a possible confound due to storage or integration effects (Gibson, 2000). The remaining sentences should provide regions where the parser is confident about its depth estimates, where sentence length is unexceptional, and where linguistic memory load is not changing. Every third sentence is put into the exploratory development dataset, and the rest are put into the test set. For each dataset, the epochs are grouped based on their associated em-

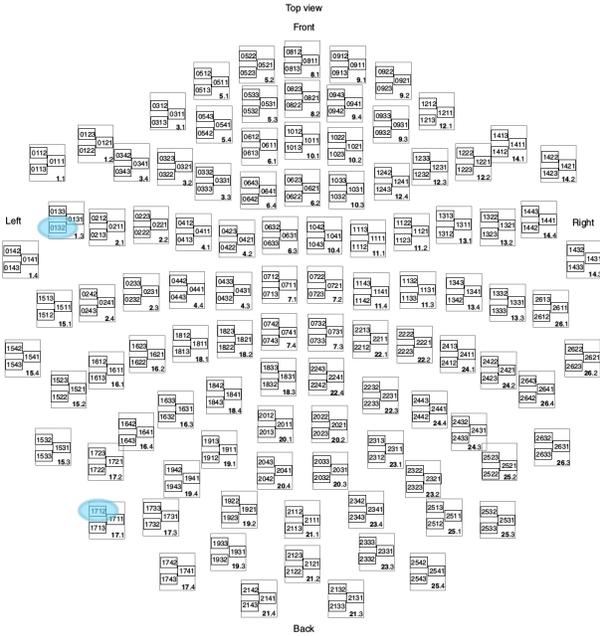


Figure 2: Top-down depiction of sensor locations in the Elekta Neuromag helmet. The front of the helmet is at the top of the figure. The sensors in blue are those used in this study.

bedding depths. Each depth grouping is further clustered into sets of four epochs; these sets are used to calculate the expectations necessary to compute coherence.<sup>7</sup> Continuous wavelet decomposition (Gabor, 1946) is employed to decompose the waveform signal recorded by each sensor into its component frequencies.

The memory load of a given epoch should be relatively constant throughout the duration of a given word, so the dependent variable tested in this study is the average coherence from 0-500 ms after the onset of each word. If the average coherence of a frequency is high due to a brief spike in coherence during that window rather than due to repeated synchronous firing of the neural clusters under investigation, the increased variance will penalize the sig-

<sup>7</sup>The choice to cluster into sets of four epochs was driven by the data. In order to obtain valid statistical significance in the development data regarding embeddings at an embedding depth of one, the data could only be divided by 4 before  $n$  dropped below 30. While statistical significance is not needed for exploration, a less-than-representative sample in the development set would negate the purpose of having a development set for exploration.

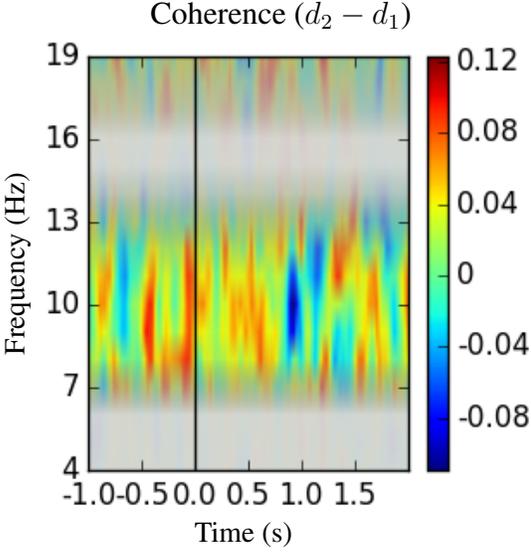


Figure 3: A time-frequency depiction of the mean coherence in the depth = 1 condition subtracted from the mean coherence in the depth = 2 condition in development data. An overlay conveys the variance of different frequency bands. Faded regions have higher variance than clearer regions.

nificance of that frequency. Although this work initially averaged over epochs during computation of coherence in order to obtain good frequency resolution, exploration using development data revealed that coherence often appears across several adjacent frequency bands, so to boost the signal-to-noise ratio, the dependent variable was recast as the average coherence within  $\pm 2$  Hz of each frequency band. Since this study is focused on linguistic processing, the development data was searched for two sensors in the anterior and posterior regions of the left hemisphere with a high degree of depth-sensitive alpha coherence. In analysis of the development set, gradiometer sensors 0132 and 1712 (anterior and posterior sensors, respectively; shown in Figure 2) showed a high coherence, so these were used in the evaluation on the test set. This was the only sensor pair evaluated on the test data.

To avoid making the statistical analyses vulnerable to assumptions about data distribution, statistical significance of depth as a predictor in the development and test datasets is calculated using the Mann-Whitney  $U$ -test, a non-parametric alternative to the  $t$ -test for testing differences between two unpaired

Factor	Coef	p-value
Unigram	$5.1 \cdot 10^{-5}$	0.941
Bigram	$5.6 \cdot 10^{-4}$	0.257
Trigram	$4.3 \cdot 10^{-4}$	0.073
PCFG Surprisal	$2.8 \cdot 10^{-4}$	0.482
Sentence Position	$-5.1 \cdot 10^{-4}$	0.031
Depth	$3.6 \cdot 10^{-2}$	0.005

Table 1: Development data results using each factor to predict alpha coherence from 0-500ms at  $10 \pm 2$ Hz.

samples. The  $U$ -test is used to see whether the distribution of coherence at a given depth is the same as the distribution of coherence at another depth.

Development analysis finds that the depth 1 data ( $n = 40$ ) and the depth 2 data ( $n = 1118$ ) have significantly different coherence distributions around 10 Hz ( $p = 0.005$ ; see Figure 3), which is in the middle of the alpha frequency range (8-12Hz). This finding suggests that alpha coherence between these two regions are predictive of linguistic working memory load. To ensure that this finding was not caused by a single subject, the same analysis was repeated over the development data after omitting each subject in turn, with similar results.

It may be, however, that these alpha coherence effects are driven by confounding factors like sentence position (alpha coherence may be more likely to occur near the beginnings or ends of sentences) or frequency (alpha coherence may tend to increase when processing rare or common words), which may be collinear with depth. In order to check for these possible confounds, the data must be re-ordered by sentence position or frequency predictors, then re-grouped into sets of four before computing coherence, in order to avoid computing coherence over unrelated factor levels.<sup>8</sup>

To rule out the confounds of sentence position and frequency, a variety of independent predictors are separately linearly regressed against the dependent variable of coherence. Four different frequency predictors are used: unigrams, bigrams, trigrams, and PCFG surprisal. The  $n$ -gram factors are all log-probabilities computed from the Corpus of Contem-

<sup>8</sup>Since only two values of depth are tested in the present study, depth is always tested using a  $U$ -test, while the more continuous variables are tested using linear regression.

Factor	Coef	p-value
Unigram	$-2.2 \cdot 10^{-4}$	0.6480
Bigram	$-9.8 \cdot 10^{-5}$	0.7762
Trigram	$3.7 \cdot 10^{-4}$	0.0264
PCFG Surprisal	$2.9 \cdot 10^{-4}$	0.3295
Sentence Position	$1.3 \cdot 10^{-4}$	0.4628
Depth	$4.6 \cdot 10^{-2}$	0.00002

Table 2: Test data results using each factor to predict alpha coherence from 0-500ms at  $10 \pm 2$ Hz. Note that the trigram factor is not a significant predictor after applying Bonferroni correction.

porary American English (COCA; Davies, 2008) and PCFG surprisal is computed by the van Schijndel et al. (2013) incremental parser. While sentence position is significant on the development partition (Table 1), none of the frequency-based effects are significant in the development set, but this may be due to having too little data in the development set, so all factors are tested again in the larger test set.<sup>9</sup>

To retain an  $\alpha$ -level of 0.05 with six statistical tests, the threshold for significance must be Bonferroni corrected to 0.008. As shown in Table 2, sentence position fails to be a significant predictor of alpha coherence on the test data (even without Bonferroni correction), but embedding depth remains a significant predictor of alpha coherence. The marginal effect of trigram predictability observed in the development set remains in the test set, but the effect is not significant after correcting for multiple comparisons.

While Bonferroni correction would rule out trigram probability as a significant predictor even if it was the only non-depth predictor tested in this work, the fact that it is marginally significant in both datasets is suggestive of a true underlying effect. To determine whether trigram probability is actually predictive of MEG coherence, we increase the resolution of the coherence by using six epochs (rather than the previous four) to compute the expectations in Equation 1. The increase in resolution further

<sup>9</sup>In development testing, ‘significance’ is merely a convenient tool for summarizing how strongly correlated the independent and dependent variables are. The general lack of correlation between MEG coherence and position/frequency predictors in development data suggests this is a promising dependent variable for our purposes.

Factor	Coef	p-value
Trigram	$1.6 \cdot 10^{-4}$	0.3817
Depth	$3.2 \cdot 10^{-2}$	0.0046

Table 3: Test data results after increasing coherence resolution to six epochs.

shrinks the dataset, but the larger test set can absorb the loss and still provide valid significance results.<sup>10</sup> The results (Table 3) show that, with greater coherence resolution, embedding depth remains a significant predictor of MEG coherence, and that trigram probability is not even a marginally significant predictor. These results reinforce the theory that alpha coherence reflects memory load and further shows that alpha coherence between the anterior and posterior regions of the left hemisphere may specifically reflect linguistic memory load.

## 7 Discussion

This study found that alpha coherence between the anterior and posterior regions of the left hemisphere of the brain is significantly correlated with embedding depth, which suggests that alpha coherence may reflect an effect of memory load on linguistic processing in those regions. This correlation was found in an exploratory study using development data and subsequently confirmed by generalizing to held-out test data. These results are consistent with patterns observed in fMRI experiments: a large survey (Hagoort and Indefrey, 2014) identifies activation of the left inferior frontal gyrus (LIFG, including “Broca’s area”) and posterior parts of the left temporal cortex (including “Wernicke’s area”), during both passive listening and passive reading tasks. Their findings indicate that, with listening tasks in particular, the anterior region of the right hemisphere is also active, and the results of Weiss et al. (2005) suggest that EEG coherence between the left and right hemispheres of the brain increases with embedding depth. Future study is needed to determine if rightside coherence or left-right coherence in MEG data is also associated with embedding depth.

Importantly, the alpha coherence found in this

<sup>10</sup>After increasing coherence resolution, trigram  $n = 1933$ , depth 1  $n = 57$ , and depth 2  $n = 1428$ .

study did not correlate with sentence position or frequency effects. The lack of influence of position and frequency effects on MEG coherence could greatly facilitate future research on sentence processing, since these effects often present large confounds in predicting other psycholinguistic measures. The cost associated with collecting MEG data may limit the immediate widespread application of the present findings, but since MEG and EEG signals are produced by electrical activity from the same underlying brain sources, this gives hope that anterior-posterior left hemisphere alpha coherence in EEG may be able to provide a similarly clear signal for future studies.

The present data support findings like those of van Schijndel and Schuler (2015), who claim hierarchic structure must be used during linguistic processing because hierarchic structure improves the fit to reading times over competitive non-hierarchic models. A potential criticism of that finding is that humans may make use of linear sequences of part-of-speech tags but not hierarchic structure during linguistic processing (Frank and Bod, 2011). In that case, the improved fit of the hierarchic grammars in van Schijndel and Schuler (2015) may simply stem from the fact that hierarchic grammars also happen to contain part-of-speech information as well as hierarchic structure. The findings of the present study support the theory that hierarchic structure is used during linguistic processing since this study finds a clear effect of alpha coherence conditioned on hierarchic embedding depth.

Having identified a working-memory based signal that is seemingly free of many of the confounding influences associated with reading times, it should be interesting to use the same procedure to study linguistic regions where embedding depth changes. Such studies could tell us what activation patterns arise due to storage and integration of linguistic elements in working memory. Contrary to the previous studies of such influences, which relied on indirect measures such as reading time latencies, if coherence is construed as attentional focus (Jensen et al., 2007), the present methods could directly investigate theoretical claims such as those made by Gibson (2000) and Lewis et al. (2006) regarding the attentional resources required for storage and integration of incomplete dependencies under different

conditions. That is, it permits direct measurement of whether and how much attentional resources must be expended in cohering disparate regions of the brain in those conditions. Such resource expenditures could manifest themselves in reading times in a variety of ways, but the present work has outlined a technique, seemingly independent of frequency effects, of directly testing the underlying theoretical linguistic claims in naturalistic data.

## Acknowledgements

Thanks to the anonymous reviewers for their helpful feedback on this work. Thanks also to Roberto Zamparelli of the University of Trento, for advice on the pilot experimental and analytical design. This material is based upon work supported by the National Science Foundation Graduate Research Fellowship under Grant No. DGE-1343012. Support also came from the University of Pittsburgh Medical Center MEG Seed Fund, and the National Institutes of Health CRCNS grant No. 5R01HD075328-02. Any opinion, findings, and conclusions or recommendations expressed in this material are those of the authors and do not necessarily reflect the views of the funders.

## References

Steven P. Abney and Mark Johnson. 1991. Memory requirements and local ambiguities of parsing strategies. *J. Psycholinguistic Research*, 20(3):233–250.

Alfred V. Aho and Jeffery D. Ullman. 1972. *The Theory of Parsing, Translation and Compiling; Volume. I: Parsing*. Prentice-Hall, Englewood Cliffs, New Jersey.

V. A. Benignus. 1969. Estimation of the coherence spectrum and its confidence interval using the fast Fourier transform. *IEEE Transactions on Audio and Electroacoustics*, 17(2):145–150.

Jonathan R Brennan. 2010. *Incrementally Dissociating Syntax and Semantics*. Ph.D. thesis, New York University.

Noam Chomsky and George A. Miller. 1963. Introduction to the formal analysis of natural languages. In *Handbook of Mathematical Psychology*, pages 269–321. Wiley, New York, NY.

Mark Davies. 2008. The corpus of contemporary american english: 450 million words, 1990-present.

Arnaud Delorme and Scott Makeig. 2003. EEGLAB: an open source toolbox for analysis of single-trial dynam-

ics including independent component analysis. *Journal of Neuroscience Methods*, 134(1):9–21, mar.

Vera Demberg and Frank Keller. 2008. Data from eye-tracking corpora as evidence for theories of syntactic processing complexity. *Cognition*, 109(2):193–210.

Juergen Fell and Nikolai Axmacher. 2011. The role of phase synchronization in memory processes. *Nature Reviews Neuroscience*, 12(2):105–118.

Stefan Frank and Rens Bod. 2011. Insensitivity of the human sentence-processing system to hierarchical structure. *Psychological Science*.

Dennis Gabor. 1946. Theory of communication. *Journal of the IEEE*, 93:429–441.

Edward Gibson. 1991. *A computational theory of human linguistic processing: Memory limitations and processing breakdown*. Ph.D. thesis, Carnegie Mellon.

Edward Gibson. 2000. The dependency locality theory: A distance-based theory of linguistic complexity. In *Image, language, brain: Papers from the first mind articulation project symposium*, pages 95–126, Cambridge, MA. MIT Press.

A. Gramfort, M. Luessi, E. Larson, D. Engemann, D. Strohmeier, C. Brodbeck, R. Goj, M. Jas, T. Brooks, L. Parkkonen, and M. Hämäläinen. 2013. MEG and EEG data analysis with MNE-python. *Frontiers in Neuroscience*, 7:267.

A. Gramfort, M. Luessi, E. Larson, D. Engemann, D. Strohmeier, C. Brodbeck, L. Parkkonen, and M. Hämäläinen. 2014. MNE software for processing MEG and EEG data. *NeuroImage*, 86:446–460.

Peter Hagoort and Peter Indefrey. 2014. The neurobiology of language beyond single words. *Annual Review of Neuroscience*, 37:347–362.

Peter Hagoort. 2008. The fractionation of spoken language understanding by measuring electrical and magnetic brain signals. *Philosophical Transactions of the Royal Society B: Biological Sciences*, 363(1493):1055–1069, mar.

John Hale. 2001. A probabilistic early parser as a psycholinguistic model. In *Proceedings of the second meeting of the North American chapter of the Association for Computational Linguistics*, pages 159–166, Pittsburgh, PA.

Ole Jensen, Jack Gelfand, John Kounios, and John E. Lisman. 2002. Oscillations in the alpha band (9–12 hz) increase with memory load during retention in a short-term memory task. *Cerebral Cortex*.

Ole Jensen, Jochen Kaiser, and Jean-Philippe Lachaux. 2007. Human gamma-frequency oscillations associated with attention and memory. *Trends in Neurosciences*, 30(7):317–324.

Philip N. Johnson-Laird. 1983. *Mental models: towards a cognitive science of language, inference, and*

- consciousness*. Harvard University Press, Cambridge, MA, USA.
- Roger Levy. 2008. Expectation-based syntactic comprehension. *Cognition*, 106(3):1126–1177.
- Richard L. Lewis and Shravan Vasishth. 2005. An activation-based model of sentence processing as skilled memory retrieval. *Cognitive Science*, 29(3):375–419.
- Richard L. Lewis, Shravan Vasishth, and Julie A. Van Dyke. 2006. Computational principles of working memory in sentence comprehension. *Trends in Cognitive Science*, 10(10):447–454.
- Mitchell P. Marcus, Beatrice Santorini, and Mary Ann Marcinkiewicz. 1993. Building a large annotated corpus of English: the Penn Treebank. *Computational Linguistics*, 19(2):313–330.
- Luan Nguyen, Marten van Schijndel, and William Schuler. 2012. Accurate unbounded dependency recovery using generalized categorial grammars. In *Proceedings of the 24th International Conference on Computational Linguistics (COLING '12)*, pages 2125–2140, Mumbai, India.
- Slav Petrov and Dan Klein. 2007. Improved inference for unlexicalized parsing. In *Proceedings of NAACL HLT 2007*, pages 404–411, Rochester, New York, April. Association for Computational Linguistics.
- Liina Pykkänen and Alec Marantz. 2003. Tracking the time course of word recognition with MEG. *Trends in cognitive sciences*, 7(5):187–189.
- Philip Resnik. 1992. Left-corner parsing and psychological plausibility. In *Proceedings of COLING*, pages 191–197, Nantes, France.
- Brian Roark, Asaf Bachrach, Carlos Cardenas, and Christophe Pallier. 2009. Deriving lexical and syntactic expectation-based measures for psycholinguistic modeling via incremental top-down parsing. *Proceedings of the 2009 Conference on Empirical Methods in Natural Language Processing*, pages 324–333.
- Riitta Salmelin. 2007. Clinical neurophysiology of language: the meg approach. *Clinical Neurophysiology*, 118(2):237–254.
- William Schuler and Marten van Schijndel. 2014. Effects of integration in eye tracking. In *Twenty-Seventh Annual CUNY Conference on Human Sentence Processing*, page 207.
- Per B Sederberg, Michael J Kahana, Marc W Howard, Elizabeth J Donner, and Joseph R Madsen. 2003. Theta and gamma oscillations during encoding predict subsequent recall. *The Journal of Neuroscience*, 23(34):10809–10814.
- Elisabet Service, Päivi Helenius, Sini Maury, and Riitta Salmelin. 2007. Localization of Syntactic and Semantic Brain Responses using Magnetoencephalography. *Journal of Cognitive Neuroscience*, 19(7):1193–1205.
- Wolf Singer. 1999. Neuronal synchrony: A versatile code for the definition of relations? *Neuron*, 24(1):49–65.
- Nathaniel J. Smith and Roger Levy. 2013. The effect of word predictability on reading time is logarithmic. *Cognition*, 128:302–319.
- Edward Stabler. 1994. The finite connectivity of linguistic structure. In *Perspectives on Sentence Processing*, pages 303–336. Lawrence Erlbaum.
- Samu Taulu and Riitta Hari. 2009. Removal of magnetoencephalographic artifacts with temporal signal-space separation: demonstration with single-trial auditory-evoked responses. *Human Brain Mapping*, 30:1524–1534.
- Marten van Schijndel and William Schuler. 2015. Hierarchic syntax improves reading time prediction. In *Proceedings of NAACL-HLT 2015*. Association for Computational Linguistics.
- Marten van Schijndel, Andy Exley, and William Schuler. 2013. A model of language processing as hierarchic sequential prediction. *Topics in Cognitive Science*, 5(3):522–540.
- Marten van Schijndel, William Schuler, and Peter W Culicover. 2014. Frequency effects in the processing of unbounded dependencies. In *Proc. of CogSci 2014*. Cognitive Science Society.
- Shravan Vasishth and Richard L. Lewis. 2006. Argument-head distance and processing complexity: Explaining both locality and antilocality effects. *Language*, 82(4):767–794.
- Christoph von der Malsburg. 1995. Binding in models of perception and brain function. In *Current Opinion in Neurobiology*, pages 520–526.
- Leila Wehbe, Brian Murphy, Partha Talukdar, Alona Fyshe, Aaditya Ramdas, and Tom Mitchell. 2014a. Simultaneously uncovering the patterns of brain regions involved in different story reading subprocesses. *PLoS one*, 9(11):e112575.
- Leila Wehbe, Ashish Vaswani, Kevin Knight, and Tom Mitchell. 2014b. Aligning context-based statistical models of language with brain activity during reading. In *Proceedings of the 2014 Conference on Empirical Methods in Natural Language Processing*. ACL.
- Sabine Weiss, Horst M. Mueller, Baerbel Schack, Jonathan W. King, Martha Kutas, and Peter Rappelsberger. 2005. Increased neuronal communication accompanying sentence comprehension. *International Journal of Psychophysiology*, 57:129–141.
- P.C. Woodland, J.J. Odell, V. Valtchev, and S.J. Young. 1994. Large vocabulary continuous speech recognition using HTK. In *Proceedings of ICASSP '94*.

*IEEE International Conference on Acoustics, Speech and Signal Processing*, volume ii, pages II/125–II/128. IEEE.

Stephen Wu, Asaf Bachrach, Carlos Cardenas, and William Schuler. 2010. Complexity metrics in an incremental right-corner parser. In *Proceedings of the 48th Annual Meeting of the Association for Computational Linguistics (ACL'10)*, pages 1189–1198.